# Desirability Function Approach on the Optimization of Multiple Bernoulli-distributed Response

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Abstract: The multiple response optimization (MRO) problem is commonly found in industry and many other scientific areas. During the optimization stage, the desirability function method, first proposed by Harrington (1965), has been widely used for optimizing multiple responses simultaneously. However, the formulation of traditional desirability functions breaks down when the responses are Bernoulli-distributed. This paper proposes a simple solution to avoid this breakdown. Instead of the original binary responses, their probabilities of defined outcomes are considered in the logistic regression models and they are transformed into the desirability functions. An example is used for demonstration.

### **1 INTRODUCTION**

As science and technology have advanced to a higher level nowadays, investigators are becoming more interested in and capable of studying large-scale systems. In industry, engineering and many other areas of science, data collected often contain several responses of interest for a single set of explanatory variables. There are plenty of model selection methods, like the LASSO (Tibshirani, 1996), Dantzig selector (Candes and Tao 2006; Phoa et al., 2009), SRRS (Phoa, 2012ab) and so on, to find a setting of the explanatory variables that optimizes a single response. However, when multiple responses are required to be optimized simultaneously, it is usually difficult to come up with an optimal setting, or even several feasible settings, of explanatory variables.

Multiple response problems (Khuri, 1996; Kim and Lin, 2006) consists of three stages: data collection (design of experiments), model building, and optimization, specifically called multiple response optimization (MRO). There exists several popular approaches to reduce multiple responses to one with a single aggregated measure and solves it as a single objective optimization problem. They include the desirability function (Harrington, 1965; Derringer and Suich, 1980; Kim and Lin, 2000), the generalized distance measure method (Khuri and Conlon, 1981), a square error loss function (Pignatiello, 1993; Vining, 1998), a goal attainment approach (Xu et al., 2004), and so on. Simple linear regression is often used to investigate the relationship between a single explanatory variable and a single response, but often the response is not a numerical value. Instead, the response is simply a designation of one of two possible outcomes, e.g. yes or no, accept or decline, etc. In fact, data involving the relationship between explanatory variables and Bernoulli-distributed responses abound in just about every discipline from engineering to, the natural sciences, medicine, education, etc. Thus, it becomes a challenge to deal with the optimization of multiple responses such that these responses are binary.

The goal of this paper is to propose a modified method to the Harrington's desirability function approach to adapt with the optimization of multiple Bernoulli-distributed responses. In section 2, the Harrington's desirability function is introduced and a discussion is included on how the formulation is broken down when binary responses are dealt. In section 3, a modified method is proposed on the model building procedure prior to the desirability function, and the formulation becomes more simplified. Section 4 provides an example to demonstrate how the proposed method works, and some concluding remarks are included in the last section.

## 2 THE DESIRABILITY FUNCTION AND ITS LIMITATION TO BINARY RESPONSES

The desirability function method transforms each response into a dimensionless individual desirabilities and then combines these individual desirabilities into one whole desirability using a geometric mean. Generally speaking, when an experimental result with *m* responses  $y = (y_1, \ldots, y_m)$  and *k* factors  $x = (x_1, \ldots, x_k)$  is given, *m* models can be built for each response using a common model selection approach, and this leads to *m* fitted responses  $\hat{y} = (\hat{y}_1, \ldots, \hat{y}_m)$ . Then each fitted response  $\hat{y}_i$  is transformed into an individual desirability value  $d_i$ ,  $0 \le d_i \le 1$ . The overall desirability, denoted by *D*, is the geometric mean of all the transformed responses, given by

$$D = (d_1 \times \cdots \times d_m)^{1/n}$$

The value of  $d_i$  increases as the desirability of the corresponding response increases. The single value of D gives the overall assessment on how desirable is the combined m responses under the given setting of k explanatory variables. If D is very close to 0, which means one or more individual desirabilities is close to 0, then the corresponding setting would not be acceptable. On the other hands, if D is close to 1, then all of the individual desirabilities are simultaneously close to 1, thus the corresponding setting would be a good compromise among the m responses. The optimization goal in this method is to find the maximum of the overall desirabilities D and its associated optimal setting of explanatory variables.

The transformation from  $\hat{y}_i$  to  $d_i$  can be either onesided or two-sided. One-sided transformations are used when the goal is to either maximize or minimize the response, while two-sided transformations are used when the goal is for the response to achieve some specified target value. Harrington (1965) used exponential functions to transform  $\hat{y}_i$  to  $d_i$ , specifically  $d_i = exp(-exp(-\hat{y}_i))$  for a one-side transformation and  $d_i = exp(-\|\hat{y}_i\|^r)$  for a two-sided transformation, where r is a user-selected shape parameter that should be carefully chosen to reflect expert opinion. Derringer and Suich (1980) modified Harrington's transformations and classified them into three forms. When the goal is to maximize the *i*<sup>th</sup> response, the individual desirability is given by the one-sided transformation

$$d_{i} = \begin{cases} 0, & \hat{y}_{i}(x) < L \\ (\frac{\hat{y}_{i}-L}{U-L})^{r}, & L \leq \hat{y}_{i} \leq U \\ 1, & \hat{y}_{i}(x) > U \end{cases}$$

where U and L are acceptable maximum and minimum values of the response  $y_i$  respectively, and r is a user-specified weight describing the shape of the desirability function. Similarly, to minimize the  $i^{th}$  response, the individual desirability is given by the onesided transformation

$$d_{i} = \begin{cases} 1, & \hat{y}_{i}(x) < L \\ (\frac{U - \hat{y}_{i}}{U - L})^{r}, & L \leq \hat{y}_{i} \leq U \\ 0, & \hat{y}_{i}(x) > U \end{cases}$$

When the goal is to obtain a target value, the individual desirability is given by the two-sided transformation

$$d_{i} = \begin{cases} 0, & \hat{y}_{i}(x) < L \\ (\frac{\hat{y}_{i}-L}{T-L})^{r_{1}}, & L \leq \hat{y}_{i} \leq T \\ (\frac{U-\hat{y}_{i}}{U-T})^{r_{2}}, & T \leq \hat{y}_{i} \leq U \\ 0, & \hat{y}_{i}(x) > U \end{cases}$$

where *T* is the target value of the response  $y_i$ ,  $r_1$  and  $r_2$  are user-specified weights describing the shapes of two-sided desirability function. Derringer (1994) propose an extended an general form of *D*, using a weighted geometric mean, given by

$$D = (d_1^{w_1}, \ldots, d_m^{w_m})^{1/\sum_i w_i}$$

where  $w_i$  is the  $i^{th}$  weight on the  $i^{th}$  response specified by users.

The desirability function approach works fine when the responses are continuous. However, when the responses are Bernoulli-distributed, the ordinary regression model does not provide meaningful fitted responses. Let's consider a simple example. Given a binary response  $y_1$  where +1 and -1 correspond to YES and NO respectively, and let a setting of x return a fitted value  $\hat{y}_1 = 0.8$  through a linear regression model. If the goal is to maximize the response, following the traditional desirability function approach, a upper and a lower bound of  $y_1$  has to be found prior to the transformation of  $d_1$  from  $y_1$  under a setting of x. Let's say the upper bound U = 0.9 and the lower bound L = 0.5 in the fitted response  $\hat{y}_1$ , and set r = 1, then  $d_1 = (0.8 - 0.5)/(0.9 - 0.5) = 0.75$ . Although it is mathematically possible to compute the individual desirability  $d_1$ , it is very difficult to interpret its meaning because neither  $\hat{y}_1$ , U nor L carry any meanings that correspond to YES or NO, and thus the arithmetics among them seem not meaningful. Therefore, instead of modeling the response via ordinary linear regression, the logistic regression is suggested in the next section.

## 3 A PROPOSED METHOD FOR MULTIPLE BERNOULLI-DISTRIBUTED RESPONSES

Unlike ordinary linear regression, logistic regression is a type of regression analysis used for predicting binary outcomes or Bernoulli trials rather than continuous outcomes. Given this difference, the logistic regression take the natural logarithm of the odds to create a continuous criterion. Mathematically speaking,

$$\log \frac{\pi}{1-\pi} = \sum_i \beta_i x_i$$

where the term in the left side is called the logit (natural logarithm of the odds). Notice that the unintuitive logit needs to be converted back to the odds via the exponential function. Therefore, although the observed variables in logistic regression are discrete, the predicted scores (logit) are modeled as a continuous variable. Notice that  $\pi$  and  $1 - \pi$  are the probabilities that the outcomes are +1 and -1 respectively.

By some simple algebra, the fitted probability  $\pi$  (the outcome is +1) can be rewritten as

$$\pi(x) = \frac{1}{1 + e^{-\sum_i \beta_i x_i}}$$

Since these fitted probabilities are continuous, they can serve as the substitutions of the Bernoullidistributed responses to transform into the desirability function. In general, the proposed method follows the steps below. Given an experimental result that consists of k continuous factors x and m Bernoullidistributed responses y,

- 1. Fit *m* logistic regression models with  $x_1, \ldots, x_k$  and the corresponding estimates  $\beta_1, \ldots, \beta_k$  are obtained, where  $\beta_i$  is a vector of length *m*.
- 2. For each setting of x in a trial, obtain m fitted probabilities  $\pi_1, \ldots, \pi_m$ .
- 3. Transform the fitted probabilities  $\pi_1, \ldots, \pi_m$  into individual desirability function  $d_1, \ldots, d_m$ .
- 4. Obtain the overall desirability function *D*, which is the geometric mean of the individual desirability functions.

Due to the nature of the fitted probabilities, the optimization goal is to either maximize or minimize  $\pi$ . When the goal is to maximize the *i*<sup>th</sup>  $\pi$ , in other words, to maximize the probability of *i*<sup>th</sup> response as +1 (with a defined meaning like YES or accepted),

the individual desirabilities is given by

$$d_i = \begin{cases} 0, & \hat{\pi}_i(x) < L\\ (\frac{\hat{\pi}_i - L}{U - L})^r, & L \leq \hat{\pi}_i \leq U\\ 1, & \hat{\pi}_i(x) > U \end{cases}$$

where U and L are acceptable maximum and minimum probabilities that  $i^{th}$  response as +1, and r is a user-specified weight. It is obvious that  $U \le 1$  and  $L \ge 0$ . If both equalities hold, the individual desirability function can be simplified as  $d_i = \hat{\pi}_i$ , or simply the probability of  $i^{th}$  response as +1. Similarly, when the goal is to minimize the  $i^{th} \pi$ , in other words, to minimize the probability of  $i^{th}$  response as +1, or equivalently, to maximize the probability of  $i^{th}$  response as -1 (with a defined meaning like NO or rejected), the individual desirabilities is given by

$$d_i = \begin{cases} 1, & \hat{\pi}_i(x) < L \\ (\frac{U - \hat{\pi}_i}{U - L})^r, & L \le \hat{\pi}_i \le U \\ 0, & \hat{\pi}_i(x) > U \end{cases}$$

It is obvious again that  $U \le 1$  and  $L \ge 0$ . If both equalities hold, the individual desirability function can be simplified as  $d_i = 1 - \hat{\pi}_i$ , or simply the probability of  $i^{th}$  response as -1. Notice that it is not necessary to define the two-sided transformation because a target value *T* in between 0 and 1 is meaningless to the optimization process in Bernoulli-distributed responses. For example, if a target value T = 0.4 is desired for a particular response, it means the optimized response is a linear combination of both +1 and -1 with some weights. This violates the nature of the response that there are only two possible choices (+1 and -1).

#### **4** AN ILLUSTRATIVE EXAMPLE

Vander Heyden et al., (1999) used the highperformance liquid chromatography (HPLC) method to study the assay of ridogrel and its related compounds in ridogrel oral film-coated tablet simulations. They chose to use a 12-run Plackett-Burman design to identify the importance of eight factors on seven responses. We consider only two out of seven specific responses, which are the percentage recovery of ridogrel (%*MC*) and analysis time  $(t_R)$ . Both responses are continuous variables. The Stepwise Response Refinement Screener (SRRS) proposed by Phoa (2012a) identifies that Factors E and F (the percentage organic solvent in the mobile phase at the start and at the end of the gradient) have significant impact to %MC, and Factors B (Column Manufacturer) and F have significant impact to  $t_R$ . The ordinary linear regression models of %MC and  $t_R$  have p-values less than 0.0005.

Table 1 gives three factors, design matrix and the observed two responses. There are two forms of observed responses: continuous and binary. The continuous responses are the original data from Vander Heyden et al. (1999). However, this example aims at demonstrating the proposed method to deal with multiple response problem, so the responses are modified into binary. In specific, the +1 label in both binary responses represent the original observed responses that are higher than their nominal values, and the -1 label represent the oposite. The logistic regression models for these two responses are

$$\hat{\pi}(\% MC) = 1/(1 + e^{-(2.4167 - 0.3333B - 0.5000E + 0.2500F)})$$
$$\hat{\pi}(t_R) = 1/(1 + e^{-(15.1667 + 1.3333B - 0.3333E - 0.1667F)})$$

where *B* is an indicator variable such that it is 1 if the column manufacturer is Prodigy, and 0 otherwise.

To obtain upper and lower bounds of  $\hat{\pi}(\% MC)$ , its logistic regression model is used on the setting of explanatory variables given in the data. Among 12 predicted responses, the maximum and minimum of them are 0.7914 and 0.3393 respectively. It is obvious that the percentage of recovery should be as high as possible, thus the one-sided transformation for maximum purpose is suggested as follows:

$$d_{1i} = egin{cases} 0, & \hat{\pi}_i(\% MC) < L \ rac{\hat{\pi}_i(\% MC) - L}{U - L}, & L \leq \hat{\pi}_i(\% MC) \leq U \ 1, & \hat{\pi}_i(\% MC) > U \end{cases}$$

where L = 0.3393 and U = 0.7914. The above desirability function suggests that it is highly undesired when  $\hat{\pi}_i(\% MC)$  is smaller than the lower bound, and it is highly recommended when  $\hat{\pi}_i(\% MC)$  is larger than the upper bound.

Table 1: High-performance Liquid Chromatography(HPLC) Experiment.

Design			Cont. R		Binary R	
B	E	F	%MC	t <sub>R</sub>	% <i>MC</i> ′	$t'_R$
Prodigy	26	45	101.6	11.500	+1	+1
Prodigy	24	41	101.7	13.000	+1	+1
Alltech	24	45	101.6	9.833	+1	$^{-1}$
Alltech	26	45	101.9	9.483	+1	-1
Alltech	24	41	101.8	10.317	+1	+1
Prodigy	24	45	101.1	12.567	+1	+1
Prodigy	24	45	101.1	12.083	+1	+1
Alltech	26	45	101.6	8.417	+1	-1
Alltech	26	41	98.4	9.200	-1	-1
Prodigy	26	41	99.7	13.800	-1	+1
Prodigy	26	41	99.7	13.317	-1	+1
Alltech	24	41	102.3	11.150	+1	+1
Norminal Value of $\% MC = 100\%$						
Norminal Value of $t_R = 9.9$ min						

To obtain upper and lower bounds of  $\hat{t}_R$ , a similar logistic regression model is used. Among 12 predicted responses, the maximum and minimum of them are 0.8410 and 0.2688 respectively. It is obvious that the analysis time should be as short as possible, thus the one-sided transformation for minimum purpose is suggested as follows:

$$d_{i} = \begin{cases} 1, & \hat{\pi}_{i}(t_{R}) < L \\ \frac{U - \hat{\pi}_{i}(t_{R})}{U - L}, & L \le \hat{\pi}_{i}(t_{R}) \le U \\ 0, & \hat{\pi}_{i}(x) > U \end{cases}$$

where L = 0.2688 and U = 0.8410. The above desirability function suggests that it is highly undesired when  $\hat{\pi}_i(t_R)$  is larger than the upper bound, and it is highly recommended when  $\hat{\pi}_i(t_R)$  is smaller than the lower bound.

### 5 SOME DISCUSSIONS AND CONCLUDING REMARKS

This paper proposes a modified method for desirability function approach to deal with the data with multiple Bernoulli-distributed responses. The proposed method avoids the breakdown on the formulation of the individual desirability function due to the difficulties on providing a meaningful explanation to the arithmetic operations.

The main modification is on the model building procedure. For each response, instead of using ordinary linear regression, the logistic regression model is suggested. Then the probability of having the original response to be +1 is transformed back from the logit of the model. This probability is then used for being transformed into the individual desirability function. Since there are only two possible choices on the original response, only one-sided transformation, either maximum or minimum, is needed to consider. An example is used for demonstrating how the proposed method works.

The popularity of the desirability function in industrial applications has not gone unnoticed and use of the desirability function is beginning to appear in other areas like clincial trials and social science. The researchese in both area contain plenty of data with multiple binary responses and the compromise setting of explanatory variables are desired. Thus the proposed method in this paper will hopefully provide a solution to these researches.

The method proposed in this paper sounds similar to some modeling approaches like multi-response logistic regression and/or penalized logistic regression. However, the main difference between them is that the regression method aims at modeling multiple responses simultaneously, but it is possible to suggest the estimates of a factor with different signs, which causes confusions when one attempts to set up the factor levels of the experiment. Desirability function is a compromise method that aggregates responses into one single quantity, and all factors are set optimally on this aggregated quantity. Thus, only one set of compromise factor setting will be returned and it reduces the confusion when one attempts to set up the experiment.

Bernoulli-distributed responses, which consists of only two possible outcomes, is the simplest case for categorical type variables. One promising direction to the next step is to develop a framework of desirability function approach for categorical responses. It is interesting to investigate in how to couple the variation information with the categorical responses when repeated experiments are done. Since these responses are not continuous, transformations on the responses and their variations are required for proper analysis. Furthermore, the example in this paper has been analyzed and thus comparable. It is desired to perform more simulations on some new real-life applications in order to check the efficiency of the generalized method for categorical responses.

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